## MSc Introductory Material - Block A

## Introduction to Probability and Statistics Exercises

- 1. Find the probabilities that a hand of four cards drawn at random from a standard pack of 52 playing cards will contain
  - (a) four cards of the same suit;
  - (b) at least two 'aces';
  - (c) the same number of 'clubs' as 'spades'.
- 2. a. Prove, using the probability axioms, that

$$P(A \cup B) = P(A) + P(B) - P(A \cap B) \tag{1}$$

b. By using the result  $(A \cup B) \cap C = (A \cap C) \cup (B \cap C)$  prove that

$$P(A \cup B \cup C) = P(A) + P(B) + P(C) - P(A \cap B) - P(A \cap C) - P(B \cap C) + P(A \cap B \cap C).$$

Write down the 'general addition law' for n events i.e.  $P(A_1 \cup A_2 \cup \ldots \cup A_n)$ .

c. Deduce from the addition law (1)

$$P(A \cup B) \leq P(A) + P(B)$$
.

Hence show,

$$P(A_1 \cup A_2 \cup ... \cup A_n) \leq P(A_1) + P(A_2) + ... + P(A_n).$$

3. a. Prove, using the result

$$(A_1 \cap A_2 \cap \ldots \cap A_n)^c = A_1^c \cup A_2^c \cup A_2^c \cup \ldots \cup A_n^c$$

and the general addition law, Bonferroni's inequality:

$$P\left(A_1\cap A_2\cap\ldots\cap A_n
ight)\geq 1-\sum_{i=1}^n P\left(A_i^c
ight).$$

b. Suppose there are n events  $A_1, A_2, \ldots, A_n$  and we require

$$P(A_1 \cap A_2 \cap \ldots \cap A_n) \ge 1 - \alpha$$

for some value lpha. What common value can be given to the individual probabilities  $P\left(A_i\right)$  to ensure that this is true?

- 4. A and B are events with 0 < P(A) < 1. Show that if  $P(B \mid A) > P(B)$ , then  $P(B \mid A^c) < P(B)$ .
- 5. If A and B are independent events, show that A and  $B^c$  are independent, and that  $A^c$  and  $B^c$  are independent. Discuss the generalization of this result to longer sequences of events.
- 6. A bag contains 5 white balls and 2 red balls. Balls are drawn at random one at a time without replacement until both red balls have been drawn. Find the probability function and distribution function of the number of draws required.
- 7. Sketch the distribution function F(x) when
  - i)  $X \sim \mathrm{Ber}ig(rac{1}{2}ig)$  (i.e.  $X \sim \mathrm{Bi}ig(1,rac{1}{2}ig)$  ),
  - ii)  $X \sim \mathrm{Bi}ig(2,rac{1}{2}ig),$
  - iii)  $X \sim Un(0,1)$  (continuous).
- 8. The geometric distribution with parameter  $\theta$  has probability function given by

$$p(x) = \left\{ egin{aligned} heta(1- heta)^{x-1} & ext{if } x=1,2,\dots \ 0 & ext{otherwise} \end{aligned} 
ight.$$

Find the mean, the second factorial moment, and hence the variance, of this distribution.

9. Verify, by direct consideration of the integrals involved, that if  $X \sim N\left(\mu, \sigma^2
ight)$ 

$$P(a \leq X \leq b) = \Phi\left(rac{b-\mu}{\sigma}
ight) - \Phi\left(rac{a-\mu}{\sigma}
ight)$$

(see 2.4.3).

- 10. Verify that the mean and variance of the Gamma distribution are as given in 2.4.4
- 11. Show that the geometric distribution and the exponential distribution have the 'lack of memory' property: that is, if X is a random variable with one of these distributions, then  $P(X > a + b \mid X > a)$  does not depend on a for a, b > 0.
- 12. If X is a random variable with  $Po(\mu)$  distribution where  $\mu$  is large, find a random variable g(X) whose approximate variance (given by the first order Taylor approximation to g) does not depend on  $\mu$ .
- 13. A r.v. X has p.d.f.  $f(x)=x/8, 0 \le x \le 4$ . Find the distribution of Y, the nearest integer to X. Compare the means and variances of X and Y.
- 14. Three counters numbered 1,2 and 3 are placed in a container. Two are drawn at random one at a time without replacement. Let X and Y denote the numbers on the two counters respectively.
  - (a) (Write down the joint probability function of ( X, Y ) and compute the probability that the number drawn first, X, is less than the number drawn second, Y.
  - (b) Obtain the marginal distributions of X and Y and the conditional distribution  $P(X=x\mid Y=y)$ .
  - (c) Are X and Y independent?
- 15. The joint p.d.f. of  $\{X_1, X_2\}$  is  $f(x_1, x_2) = ke^{-x_1}$  in the region  $R = \{(x_1, x_2) : 0 < x_2 < x_1\}$  and zero elsewhere. Find k and the marginal p.d.f. of  $X_1$ .
- 16. Verify the relation

$$\operatorname{Var}(X+Y) = \operatorname{Var} X + \operatorname{Var} Y + 2\operatorname{Cov}(X,Y)$$

given in 3.2.1.

- 17. If X has standard normal distribution, show that  $X^2$  has  $\chi^2_1$  distribution. [Check: is this transformation uniquely invertible?]
- 18. If X, Y are i.i.d. Ex(1) variables, write down the distribution of X+Y and verify this by use of this transformation U=X+Y, V=X-Y.
- 19. Suppose  $\mathbf{X} = (X_1, X_2, X_3)'$  is multivariate normal

$$\left\{ \begin{pmatrix} 3\\2\\2 \end{pmatrix}, \begin{pmatrix} 4 & -2 & -2\\-2 & 2 & 1\\-2 & 1 & 2 \end{pmatrix} \right\}$$

Obtain the joint distribution of  $Y_1=X_1+X_2, \quad Y_2=X_1+X_3$  .

20. If  $X_1, X_2, \ldots, X_n$  is a random sample from  $Ex(\lambda)$ , show that  $\bar{X}$  has sampling distribution which is  $Ga(n, n\lambda)$ . Deduce that

$$E\left(rac{1}{ar{X}}
ight) = rac{n}{n-1}\lambda \quad ext{ for } n \geq 2$$

21. Suppose that single observations  $X_1, X_2$  are taken from a Poisson  $(a\lambda)$  and a Poisson  $(b\lambda)$  respectively ( a,b are known positive constants and the observations are independent). Compare the following estimators of  $\lambda$ 

$$rac{X_1+X_2}{a+b}, \quad rac{X_1-X_2}{a-b}, \quad rac{1}{2}igg(rac{X_1}{a}+rac{X_2}{b}igg)\,.$$

22. Find the maximum likelihood estimator of  $\theta$  based on a random sample of size n from the distribution with density

$$f_X(x; heta) = heta x^{ heta-1} \quad x \in (0,1), heta > 0$$

- 23. Let X be an observation from  $\mathrm{Bi}(n,\theta)$ , where n is known and  $\theta$  is unknown. Find the maximum likelihood estimator of  $\theta$ , and show that it is unbiased.
- 24. Suppose  $X_i$  is the lifetime of patient i from time  $t_0$  and  $z_i$  is his white blood cell count at  $t_0$ . A suitable model is thought to be

$$X_{i} \sim Ex\left(\lambda_{i}
ight) ext{ where } \lambda_{i} = eta z_{i}.$$

Data is available on n patients. Find the maximum likelihood estimate of  $\beta$ .

25. The normal linear regression model is given by

$$X_i \sim N\left(lpha + eta t_i, \sigma^2
ight) ext{ for } i=1,2,\ldots,n$$

where  $t_1, t_2, \ldots, t_n$  are fixed known quantitites, and  $\alpha, \beta$  and  $\sigma^2$  are unknown parameters. Write down the log likelihood function. By differentiating it partially with respect to  $\alpha, \beta$  and  $\sigma^2$  (= v say), and setting the derivatives equal to zero, obtain expressions for the maximum likelihood estimators of  $\alpha, \beta$  and  $\sigma^2$ .

26.  $X_1,X_2,\ldots,X_n$  is a random sample from a normal distribution with mean zero and unknown variance  $\sigma^2$ . Using the test statistic  $\sum_{i=1}^n X_i^2$ , suggest the form of a test of  $H_0:\sigma^2=\sigma_0^2$  against  $H_1:\sigma^2=\sigma_1^2$ , where  $\sigma_0^2$  and  $\sigma_1^2$  are known with  $\sigma_1^2>\sigma_0^2$ . Find a test of this form with size  $\alpha=0.05$ .

If n=25, how large must  $\sigma_1^2/\sigma_0^2$  be in order for this test to have power 0.95?

Hint: use the result of q .17 and the fact that independent  $\chi^2$  variables are additive  $\left(\chi_a^2 + \chi_b^2\right) = \chi_{a+b}^2$ .

- 27. Construct a 90% confidence interval for  $\mu$  using the information that a random sample of 16 observations from a  $N\left(\mu,\sigma^2\right)$  distribution gave  $\bar{x}=95.8,s^2=30.25$ . [Hint: use the t -distribution described towards the end of 5.5.3]
- 28. Construct a 95% confidence interval for  $\sigma^2$  based on 40 observations from a  $N\left(\mu,\sigma^2\right)$  distribution which gave  $s^2=25$ .

Hint: use the result  $rac{(n-1)S^2}{\sigma^2} \sim \chi^2_{n-1}$ .

29. A single observation is taken from the  $Bi(2,\theta)$  distribution. Identify all eight possible tests of the hypotheses

$$H_0: \theta = 1/2$$
 v  $H_1: \theta = 3/4$ 

and in each case give the type I and II error probabilities. Which test would you prefer?

30. Use the Neyman-Pearson lemma to provide a test of the hypothesis  $H_0: \lambda = \lambda_0$  versus  $H_1: \lambda = \lambda_1 > \lambda_0$  based on a random sample of size n from the Poisson (  $\lambda$  ) distribution.

[Hint: use the approximation  $Po(\mu) \simeq N(\mu,\mu)$  when  $\mu$  is large to evaluate the constant.]

- 31. Find the form of the likelihood ratio test of  $H_0: \lambda = \lambda_0$  against  $H_1: \lambda \neq \lambda_0$  when  $X_1, X_2, \ldots, X_n$  is a random sample from  $Ex(\lambda)$ . Simplify it as much as possible.
- 32. Let  $X_1,X_2,\ldots,X_n$  be a random sample from  $N\left(\mu,\sigma^2\right)$ , where  $\sigma^2$  is known. Find the form of the likelihood ratio test of  $H_0:\mu=\mu_0$  against  $H_1:\mu\neq\mu_0$ . Show that  $-2\log\Lambda$  is (exactly)  $\chi_1^2$  distributed if  $H_0$  is true.

[Hint: see hint for q.26.]